

Association Between Lean Mass, Fat Mass, and Bone Mineral Density: A Meta-analysis

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Context: Body weight is the most important anthropometric determinant of bone mineral density (BMD). Body weight is mainly made up of lean mass (LM) and fat mass (FM), and which component is more important to BMD has been a controversial issue.

Objective: This study sought to compare the magnitude of association between LM, FM, and BMD by using a meta-analytic approach.

Data Source: Using an electronic and manual search, we identified 44 studies that had examined the correlation between LM, FM, and BMD between 1989 and 2013. These studies involved 20 226 men and women (4966 men and 15 260 women) aged between 18 and 92 years. We extracted the correlations between LM, FM, and BMD at the lumbar spine, femoral neck, and whole body. The synthesis of correlation coefficients was done by the random-effects meta-analysis model.

Results: The overall correlation between LM and femoral neck BMD (FNBMD) was 0.39 (95% confidence interval, 0.34 to 0.43), which was significantly higher than the correlation between FM and FNBMD (0.28; 95% confidence interval, 0.22 to 0.33). The effect of LM on FNBMD in men ($r = 0.43$) was greater than that in women ($r = 0.38$). In premenopausal women, the effect of LM on BMD was greater than the effect of FM ($r = 0.45$ vs $r = 0.30$); however, in postmenopausal women, the effects of LM and FM on BMD were comparable ($r = 0.33$ vs $r = 0.31$).

Conclusion: LM exerts a greater effect on BMD than FM in men and women combined. This finding underlines the concept that physical activity is an important component in the prevention of bone loss and osteoporosis in the population. (*J Clin Endocrinol Metab* 99: 30–38, 2014)

Low bone mineral density (BMD) is the most robust risk factor for fracture (1, 2). Each SD decrease in BMD is associated with a 2- to 3-fold increase in the risk of fracture (2), and this magnitude of association is equivalent to the association between blood pressure and cardiovascular events (3). Therefore, measurement of BMD is commonly used as a tool for the diagnosis of osteoporosis (4). Body weight is an important determinant of BMD, such that individuals with higher body weight have higher BMD (5)

and reduced fracture risk (6). Between-individual variation in body weight accounts for about 30% of variation in BMD, making it one of the best determinants of BMD (7, 8).

Body weight is largely made up of two components: fat mass (FM) and lean mass (LM; or fat-free mass). The relative contribution of each of the two components to BMD variation has been highly contentious. Although some studies have suggested that LM, not FM, is associated with

BMD (9–18), other studies (19–22) have shown that FM, not LM, is an important determinant of BMD. Still, other studies have found that both FM and LM were significant predictors of BMD (23–25). Taken together, these studies suggest that the relative contribution of body composition

parameters to BMD could be dependent on gender, ethnicity, and age. Although most studies found that LM significantly predicted BMD in both genders (26, 27), some studies observed that FM was associated with BMD for men under 50 years but not for women and men over

Table 1. Characteristics of Individual Studies Included in the Meta-Analysis

| First Author (Ref.) | Year | Ethnicity | Sex | Age, y | N | Correlation of LM | | | Correlation of FM | | |
|----------------------|------|-----------|-------------|--------|------|-------------------|------|------|-------------------|-------|-------|
| | | | | | | LS | FN | WB | LS | FN | WB |
| Bevier (53) | 1989 | Caucasian | Men | 61–84 | 36 | 0.15 | | | | | |
| | | | Women | 61–84 | 55 | 0.19 | | | | | |
| Reid (65) | 1992 | Caucasian | Men | 31 | 51 | | | 0.51 | | | 0.26 |
| | | | Women | 33 | 68 | | | 0.55 | | | 0.6 |
| Reid (20) | 1992 | Caucasian | Women | 45–78 | 140 | 0.20 | 0.18 | 0.18 | 0.34 | 0.38 | 0.55 |
| Compston (59) | 1992 | Caucasian | Women | 49–65 | 97 | 0.28 | | 0.58 | | | |
| Edelstein (68) | 1993 | Caucasian | Men | 55–84 | 597 | 0.3 | | | 0.3 | | |
| | | | Women | | 895 | 0.34 | | | 0.36 | | |
| Reid (19) | 1994 | Caucasian | Women | 47–73 | 140 | 0.21 | | | 0.43 | | |
| Reid (66) | 1995 | Caucasian | Women | 36 | 36 | 0.47 | 0.59 | 0.6 | 0.34 | 0.47 | 0.5 |
| | | | | 33 | 63 | 0.33 | 0.38 | 0.44 | 0.14 | –0.15 | 0.18 |
| Salamone (16) | 1995 | Caucasian | Women | 40–50 | 334 | 0.44 | 0.40 | 0.45 | 0.16 | 0.16 | 0.19 |
| Douchi (41) | 1997 | Asian | Women | 20–54 | 128 | 0.33 | | 0.38 | 0.38 | | 0.27 |
| | | | | | 196 | 0.47 | | 0.39 | 0.31 | | 0.23 |
| Chen (58) | 1997 | Caucasian | Women | <65 | 50 | 0.11 | 0.10 | 0.23 | | 0.07 | 0.25 |
| Barondess (55) | 1997 | Caucasian | Men | 33–64 | 42 | | | 0.58 | | | 0.35 |
| | | Black | Men | 33–64 | 37 | | | 0.51 | | | 0.60 |
| Ohmura (50) | 1997 | Asian | Women | 20–79 | 1006 | | | 0.42 | | | 0.41 |
| Nguyen (9) | 1998 | Caucasian | Women | 53 | 112 | 0.38 | 0.27 | 0.32 | 0.37 | 0.26 | 0.59 |
| Taaffe (17) | 2000 | Caucasian | Women | 60–86 | 62 | 0.34 | 0.33 | 0.29 | 0.28 | 0.31 | 0.27 |
| | | Hispanic | Women | 60–86 | 54 | 0.33 | 0.21 | 0.40 | 0.18 | 0.19 | 0.34 |
| Nakaoka (52) | 2001 | Asian | Women | 48–84 | 205 | 0.34 | 0.34 | | 0.25 | 0.38 | 0.11 |
| Lee (46) | 2001 | Asian | Women | 20–69 | 178 | 0.21 | | 0.3 | 0.09 | | 0.18 |
| Van Langendonck (67) | 2002 | Caucasian | Men | 60 | 156 | 0.33 | | 0.24 | 0.24 | | 0.20 |
| Reid (21) | 2002 | Caucasian | Women | 60 | 119 | 0.21 | | | 0.43 | | |
| Ijuin (24) | 2002 | Asian | Women | Post | 193 | 0.36 | 0.38 | 0.38 | 0.32 | 0.38 | 0.31 |
| | | Asian | Women | Pre | 360 | 0.30 | 0.39 | 0.34 | 0.04 | 0.33 | 0.09 |
| Douchi (42) | 2003 | Asian | Women | <60 | 123 | 0.31 | 0.33 | 0.32 | | | |
| | | Asian | Women | >65 | 102 | 0.38 | 0.51 | 0.42 | | | |
| Douchi (11) | 2003 | Asian | Women | | 45 | 0.42 | | | 0.15 | | |
| | | Asian | Women | | 89 | 0.23 | | | 0.25 | | |
| Douchi (10) | 2003 | Asian | Men | | 93 | 0.38 | | 0.56 | –0.03 | | –0.09 |
| Li (14) | 2004 | Caucasian | Women | 40–55 | 43 | 0.41 | 0.52 | 0.09 | 0.37 | 0.49 | 0.02 |
| Liu (15) | 2004 | Asian | Women | 20–55 | 282 | 0.28 | 0.34 | 0.36 | 0.05 | 0.14 | 0.10 |
| Wang (18) | 2005 | Mixed | Women | 20–25 | 921 | 0.38 | 0.37 | 0.41 | 0.18 | 0.23 | 0.33 |
| Mizuma (29) | 2006 | Asian | Women | 30–49 | 302 | 0.30 | | 0.35 | 0.18 | | 0.10 |
| | 2006 | Asian | Women | 50–69 | 197 | 0.42 | | 0.30 | 0.19 | | 0.20 |
| Gnudi (23) | 2007 | Caucasian | Women | 62 | 770 | | 0.33 | 0.42 | | 0.26 | 0.35 |
| Kim (45) | 2009 | Asian | Women | <50 | 1694 | 0.12 | | | 0.02 | | |
| Lee (88) | 2009 | Asian | Women | 22–72 | 60 | 0.35 | 0.47 | 0.75 | 0.25 | 0.16 | –0.04 |
| Lekamwasam (47) | 2009 | Asian | Women | 30–54 | 106 | 0.22 | 0.40 | 0.28 | 0.21 | 0.43 | 0.19 |
| Benetos (54) | 2009 | Caucasian | Men | 60–85 | 169 | 0.34 | 0.52 | 0.35 | | | |
| Ho-Pham (43) | 2010 | Caucasian | Women | 50–85 | 210 | 0.76 | 0.63 | 0.89 | 0.03 | 0.03 | 0.04 |
| Bogl (56) | 2011 | Caucasian | Women | 23–31 | 147 | 0.35 | | 0.43 | 0.46 | | 0.24 |
| | | | Men | 23–31 | 154 | 0.37 | | 0.51 | 0.43 | | 0.38 |
| Ho-Pham (44) | 2011 | Asian | Men | 18–85 | 353 | 0.45 | 0.55 | 0.47 | 0.19 | 0.16 | 0.05 |
| | | | Women | 18–85 | 863 | 0.32 | 0.32 | 0.29 | 0.09 | 0.05 | 0.01 |
| Liu (48) | 2011 | Asian | Women | 40–67 | 244 | 0.3 | 0.24 | | 0.24 | 0.25 | |
| | | | | | 298 | 0.41 | 0.25 | | 0.25 | 0.24 | |
| Dytfeld (60) | 2011 | Caucasian | Women | 52–86 | 92 | 0.23 | 0.33 | | 0.13 | 0.36 | |
| Moseley (27) | 2011 | Caucasian | Men | 40–65 | 78 | 0.16 | 0.42 | 0.42 | 0.03 | 0.11 | 0.27 |
| | | | Women | 40–65 | 56 | 0.21 | 0.38 | 0.48 | 0.20 | 0.41 | 0.57 |
| Chantler (57) | 2012 | Caucasian | Women | 18–45 | 187 | 0.09 | 0.45 | 0.53 | 0.13 | 0.27 | 0.26 |
| | | Black | Women | 18–45 | 240 | 0.33 | 0.59 | 0.41 | 0.29 | 0.53 | 0.30 |
| Park (51) | 2012 | Asian | Men + Women | 44 | 1782 | 0.12 | | 0.31 | 0.06 | | 0.20 |
| Kim (26) | 2012 | Asian | Men | >40 | 1284 | | | 0.78 | | | 0.48 |
| | | | Women | >40 | 362 | | | 0.61 | | | 0.58 |
| | | | Women | >40 | 1396 | | | 0.64 | | | 0.61 |
| El Hage (61) | 2012 | Lebanese | Men | 65–84 | 70 | | 0.48 | 0.52 | | 0.44 | 0.27 |
| Gomez-Cabello (63) | 2013 | Caucasian | Women | 65–92 | 159 | 0.41 | 0.31 | 0.47 | 0.29 | 0.25 | 0.30 |
| | | | Men | 65–92 | 64 | 0.31 | 0.11 | 0.18 | 0.01 | –0.02 | –0.12 |
| Namwongprom (49) | 2013 | Asian | Women | 40–90 | 1579 | 0.47 | 0.53 | 0.40 | 0.38 | 0.40 | 0.18 |
| Nur (64) | 2013 | Caucasian | Women | 46–75 | 202 | 0.24 | 0.26 | | 0.29 | 0.33 | |

Abbreviations: N, number of subjects; LS, lumbar spine; FN, femoral neck; WB, whole body; Pre, premenopause; Post, postmenopause.

50 years (28). It has been suggested that LM is more important than FM in premenopausal women, and FM a more important than LM in postmenopausal women (24, 28, 29), but other studies showed that only LM was associated with BMD in both premenopausal and postmenopausal women (25, 26). The inconsistent findings may be due to the strong correlation between FM and LM (9) and that body fat in Caucasian populations are generally greater than in Asian populations (30). Many past studies had relatively low sample sizes, which could contribute to the inconsistency of findings.

In the presence of conflicting findings and variability in sample sizes in individual studies, a meta-analysis may be helpful to resolve the association between body composition and bone density. Two research questions guided this study: 1) what is the real magnitude of association between body composition components and BMD; and 2) what are the effects of gender, age, and ethnicity on the association. The present study took a meta-analytical approach to address the two questions by estimating the correlation between LM, FM, and BMD.

Materials and Methods

As mentioned above, this study is a systematic review and meta-analysis that involved the synthesis of data from past studies. The study was conducted in accordance with the methods of the Cochrane Collaboration (31).

Search strategy and study inclusion

An electronic search of the literature was carried out using PubMed, Ovid, and ISI Web of Knowledge resources (all-year timespan) to identify studies relating body composition and BMD. The initial keywords used for the search included “body composition*” OR “lean mass*” OR “fat-free mass*” OR “fat mass*” concatenated with “BMD” OR “bone mass” OR “bone health.” In addition, we hand-searched review articles and checked reference lists of original articles to identify studies that might have been missed from the electronic search. The inclusion criteria were: 1) original studies published in English language journals, reporting data on body composition and BMD; 2) observational studies; 3) using dual-energy x-ray absorptiometry technology; and 4) human studies on individuals aged 18+ years. We excluded review papers, case-control and interventional studies, animal studies, and studies on children or adolescents. Two reviewers (L.T.H.-P. and T.V.N.) independently identified eligible articles according to the above criteria. Discrepancies in opinion as to whether studies should be included in the analysis were resolved by discussion.

Data extraction and synthesis

Data extraction was also done independently by two reviewers. For each study, we extracted data relating to study characteristics and outcomes. Specifically, the following data were extracted: authors, journal, year of publication, study design, ethnicity, age group, gender, number of participants, and correlation coefficient of LM/FM with BMD. If more than one paper with the same data was identified, only the one that contained the definitive data was included. The primary analysis variable was the correlation coefficient between body composition measures and BMD. The two measures were LM and FM. Three BMD sites

Table 2. Correlation Between LM, FM, and BMD: Analysis by Gender, Ethnicity, and Menopausal Status

| | Lumbar Spine BMD | Femoral Neck BMD | Whole Body BMD |
|--|------------------|-------------------|------------------|
| Overall, men and women (n = 20 226; no. of studies = 44) | | | |
| LM | 0.33 (0.29–0.36) | 0.39 (0.34–0.43) | 0.46 (0.41–0.51) |
| FM | 0.24 (0.20–0.28) | 0.28 (0.22–0.33) | 0.28 (0.21–0.31) |
| By gender | | | |
| Men (n = 4966; no. of studies = 13) | | | |
| LM | 0.36 (0.29–0.43) | 0.43 (0.27–0.60) | 0.53 (0.40–0.67) |
| FM | 0.23 (0.11–0.35) | 0.18 (0.003–0.36) | 0.23 (0.07–0.40) |
| Women (n = 15 260; no. of studies = 31) | | | |
| LM | 0.33 (0.29–0.37) | 0.38 (0.33–0.42) | 0.44 (0.39–0.49) |
| FM | 0.24 (0.20–0.29) | 0.29 (0.23–0.34) | 0.29 (0.22–0.37) |
| By ethnicity | | | |
| Asians (n = 13 730; no. of studies = 19) | | | |
| LM | 0.34 (0.29–0.39) | 0.41 (0.34–0.47) | 0.48 (0.39–0.56) |
| FM | 0.19 (0.14–0.25) | 0.29 (0.14–0.40) | 0.21 (0.11–0.31) |
| Caucasians (n = 5174; no. of studies = 23) | | | |
| LM | 0.31 (0.27–0.35) | 0.35 (0.29–0.41) | 0.44 (0.37–0.50) |
| FM | 0.29 (0.23–0.35) | 0.26 (0.18–0.33) | 0.34 (0.25–0.42) |
| By menopausal status | | | |
| Premenopause (n = 5087; no. of studies = 14) | | | |
| LM | 0.31 (0.24–0.37) | 0.45 (0.37–0.53) | 0.46 (0.39–0.52) |
| FM | 0.19 (0.10–0.27) | 0.30 (0.15–0.45) | 0.29 (0.19–0.40) |
| Postmenopause (n = 7640; no. of studies = 22) | | | |
| LM | 0.34 (0.30–0.39) | 0.33 (0.26–0.40) | 0.42 (0.34–0.50) |
| FM | 0.31 (0.26–0.35) | 0.31 (0.25–0.36) | 0.36 (0.25–0.47) |

Data are shown as coefficient of correlation (95% confidence limit).

were considered, namely, lumbar spine, femoral neck (or total hip), and whole body. Thus, a maximum of six correlation coefficients were extracted and analyzed. The correlation coefficients were initially transformed to Fisher's z-scores for meta-analysis, and then back-transformed into the original correlation coefficient in the final result.

The synthesis of z-scores across studies was done by the random-effects model (32, 33). The National Research Council 1992 (34) considers the random-effects model to be more appropriate in fitting real-world data that come from populations with varying average effect sizes with a strong assumption of representativeness. Briefly, study-level z-score (denoted by z_i) is assumed to be normally distributed with a "true" but unknown mean θ_i and a within-study variance σ^2 . The collection of θ_i across studies is assumed to follow a normal distribution with unknown mean θ and between-study variance τ^2 . The classical random-effects method recognizes that the possibility of heterogeneity of between-study variation (ie, τ^2) could be difference from zero but with a fixed value. All parameters of the random-effects model were estimated by the inverse variance weighting method as implemented by the "metafor" package (35) within the R language (36).

The heterogeneity of correlations across studies was assessed by the Cochran's Q statistic (37) and the coefficient of inconsistency (I^2). The latter is an estimate of the proportion of total variation in study estimates that is due to heterogeneity (38). Subgroup analyses by age, gender, and ethnicity were also carried out as specified in the analysis protocol. Publication bias was examined by a funnel plot (39). Furthermore, the radial plot (Galbraith plot) and the standardized residuals plot were used to assess asymmetry and publication bias (40).

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Results

Characteristics of studies

An initial search yielded 3065 articles written in English with contents relating to the body composition and BMD. However, after excluding articles that did not met the inclusion criteria, only 44 studies were included in the analysis. The 44 studies involved 20 226 individuals (15 260 women and 4966 men) with means of age ranging from 18 to 92 years. Twenty studies were conducted on Asian populations (10, 11, 15, 18, 24, 26, 29, 41–52) and 24 studies were on Caucasians (9, 14, 16, 17, 19–21, 27, 53–68) (Table 1). Thirty studies were conducted on women, five studies on men (10, 54, 55, 61, 67), and nine studies included both men and women (26, 27, 44, 51, 53, 56, 63, 65, 68).

LM, FM, and BMD

Results of random-effects analysis (Table 2) showed that the correlation between LM and BMD was greater than that between FM and BMD in men and women of all ages and ethnicities. For instance, the overall correlation between LM and BMD ranged between 0.33 (for lumbar spine BMD) to 0.46 (for whole body BMD), whereas the correlation between FM and BMD ranged between 0.24 (for lumbar spine BMD)

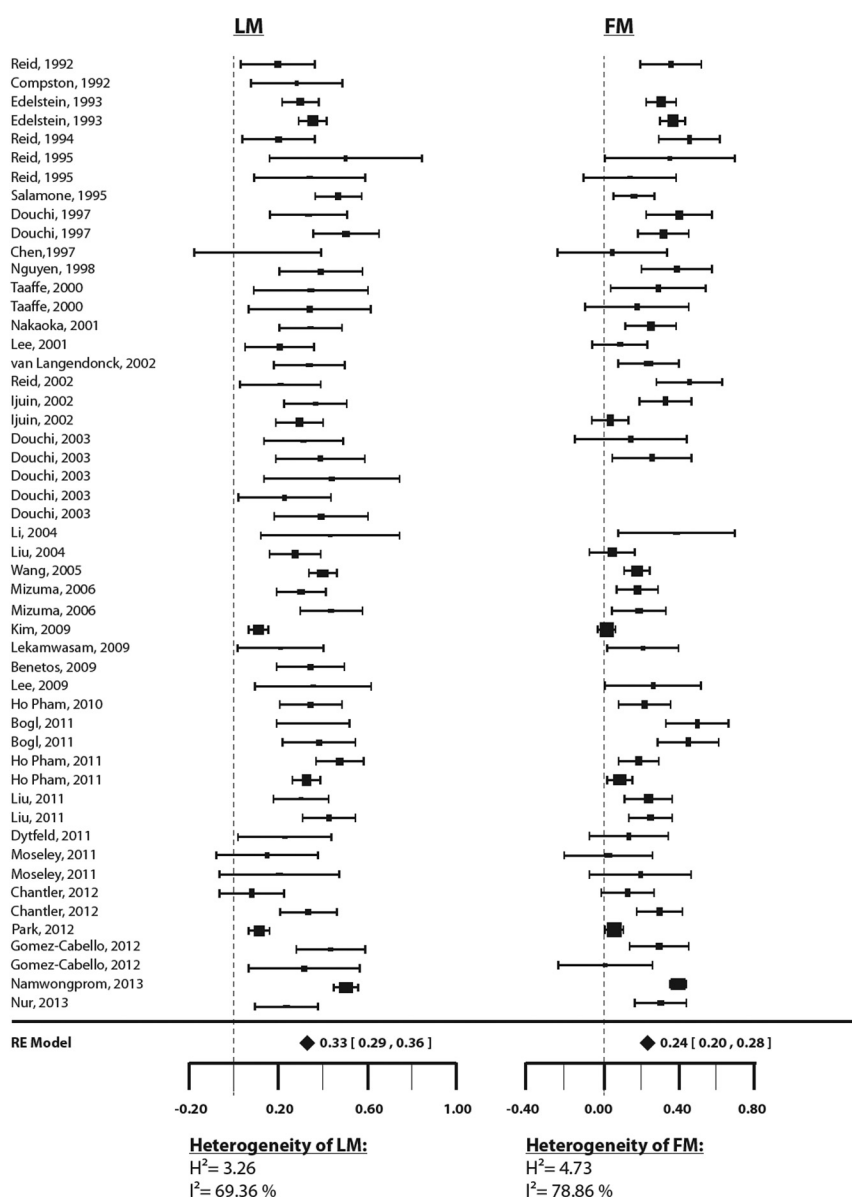


Figure 1. Correlation (and 95% confidence interval) between lumbar spine BMD and LM (left panel) and FM (right panel). The size of the dots was proportional to sample size. The overall effect size (solid diamond) was derived from the random-effects model.

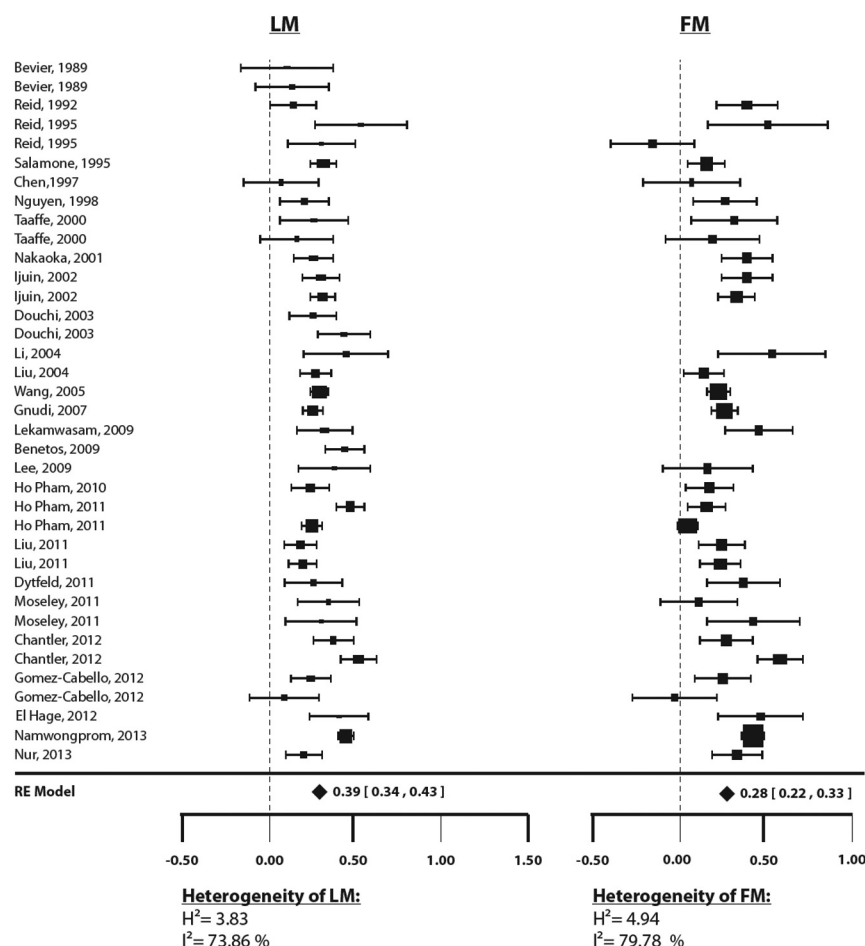


Figure 2. Correlation (and 95% confidence interval) between femoral neck BMD and LM (left panel) and FM (right panel). The size of the dots was proportional to sample size. The overall effect size (solid diamond) was derived from the random-effects model.

to 0.28 (for whole body BMD) (Figures 1–3). Measured by the correlation coefficient, it appears that the “effect” of LM on femoral neck and whole body BMD was greater than that on lumbar spine BMD. The same trend was observed in both men and women, and both Asians and Caucasians (Table 2).

Analysis by menopausal status suggested that the correlation between FM and BMD was dependent on skeletal site and menopausal status (Table 2). At the lumbar spine, the contributions of LM and FM to BMD were equivalent in postmenopausal women, but LM exerted a stronger correlation than FM in premenopausal women. For femoral neck and whole body BMD, the correlation between LM and BMD in premenopausal women ($r = 0.45$ – 0.46) was greater than that in postmenopausal women ($r = 0.33$ – 0.42).

There was a significant heterogeneity in the correlations, with the index of inconsistency (I^2) ranging from 69 to 89% for LM, and 79 to 92% for FM. However, funnel plots show no systematic trend of publication bias (Figure 4).

Discussion

For more than two decades, it has not been clear whether LM is more important than FM as a determinant of BMD. Although some studies have suggested that LM has a more pronounced effect on BMD than FM, other studies found that FM was a better determinant of BMD than LM. The discrepancy of findings is expected because previous studies have been based on different study designs with variability in sample sizes, age groups, and ethnicities. In such a context, meta-analysis offers an attractive way to resolve the issue. In this study, by using a meta-analytic approach, we have demonstrated that both LM and FM are significantly associated with BMD, but LM is more important than FM in men and women combined. We estimated that the variation in LM accounts for 21% of differences in whole body BMD, and the variation in FM explains approximately 8% of differences in BMD.

The finding that LM exerts a stronger effect on BMD than FM in men and women combined is consistent with most previous studies. In-

deed, out of 57 pairs of correlation coefficients between body composition measures and BMD, 43 coefficients showed that the correlation between LM and BMD was greater than between FM and BMD. However, the association between LM and BMD appears to be dependent on gender and age group, such that the effect of LM on BMD was stronger in men than in women, probably reflecting that the effect of muscle mass and physical activity in men is more apparent than in women. At the weight-bearing site (eg, femoral neck), the effect of LM observed in premenopausal women was greater than in postmenopausal women. In postmenopausal women, the magnitude of correlation between FM and BMD was equivalent to that between LM and BMD.

The delineation of effects of LM and FM on BMD is not easy, because LM and FM are correlated and the correlation could be different among studies due to sample size and sampling variability. If the correlation is null or low, then it is possible to estimate the contribution of each factor; if the correlation is high, the esti-

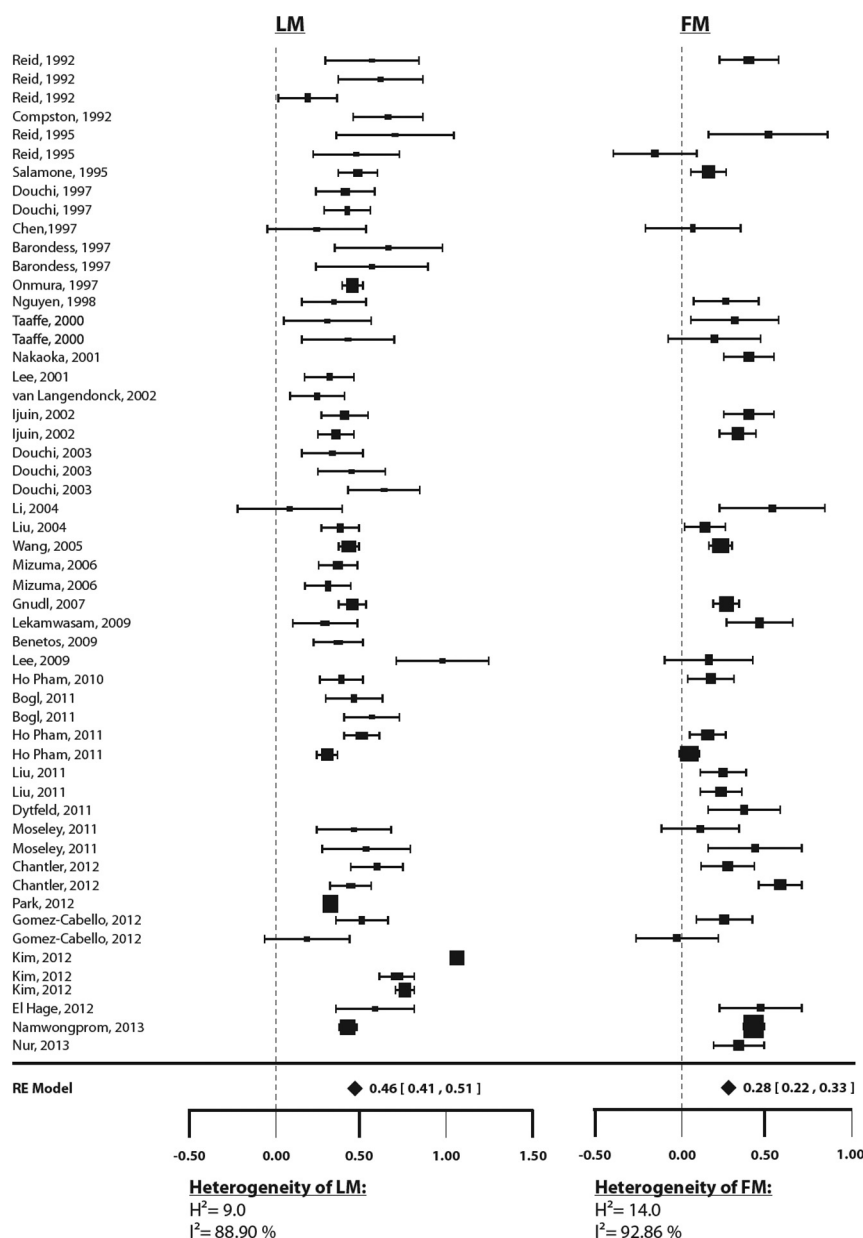


Figure 3. Correlation (and 95% confidence interval) between whole body BMD and LM (left panel) and FM (right panel). The size of the dots was proportional to sample size. The overall effect size (solid diamond) was derived from the random-effects model.

mation of contribution of any factor can be confounded by another factor. In a recent study (43), by using simulation method, it was shown that studies with sample sizes of more than 200 individuals have a much better chance of detecting the real effects of both LM and FM. These results are in broad agreement with the literature so far, in which the effects of both LM and FM were reported in studies with at least 300 individuals, and only small studies reported the effect of FM alone on BMD. In such a heuristic situation, our finding from the meta-analysis in 20 226 individuals has important implications for the identification and delineation of the effects of LM and FM on BMD.

In this study, we found that in relative terms the contribution of LM on BMD in Asians was more pronounced than in Caucasians. Based on the assumption that body fat in Caucasian populations is generally greater than in Asian populations (29), the results from some of the Caucasian studies found that FM was a strong determinant of BMD (19–21, 68), whereas most studies in Asians showed that LM was more important than FM in the association with BMD (10, 11, 15, 29, 43, 44, 49, 52). However, results from a recent study suggested that although Caucasian women have greater body weight and FM than Asian women, their percentage of body fat is similar (69), which seems to be in agreement with the finding in this study. Furthermore, it is noted that the sample size of studies on Caucasian populations was much lower than that in Asian populations. Because small sample size studies tended to yield larger and more unstable effect size than large studies, it is possible that the differential effect was purely due to sample size issue rather than biological factors.

This study also shows that there are gender-related differences in the associations between body composition and BMD. Generally, LM has a stronger relationship with BMD than FM in both genders, but this trend was more apparent in men than in women. Among women, the association between body composition

and BMD is also dependent on menopausal status. In premenopausal women, LM is more important than FM as a determinant of BMD, but in postmenopausal women, the contribution of FM to BMD variation is equivalent to that of LM.

Several hypotheses have been proposed to account for the associations between LM and FM with BMD. Both FM and LM may contribute to an increase in BMD by causing increased mechanical loading (70). In addition, the impact of LM on BMD has been attributed in part to the influence of biomechanical usage on bone development (71). According to this theory, bone strength is in-

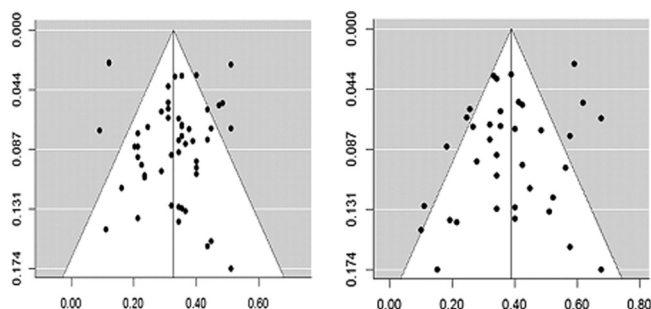


Figure 4. Funnel plot of correlation coefficients (x-axis) and SE (y-axis) for LM and femoral neck BMD (right panel) and lumbar spine BMD (left panel) vs SE.

fluenced by mechanical muscle force and hormonal factors. The force that muscle exerts against bone is influenced by how much body mass the muscles and bones support (72), and this could lead to a positive relationship between the masses of muscles and bones. As a result, physical activity seems to be important for bone mass in all periods of life and in both of genders (73), and the observed association between LM and BMD seems to be consistent with that hypothesis.

Recent data suggest that sclerostin (an osteocyte-expressed inhibitor of Wnt signaling pathway) may play a key role in the relationship between body composition and bone mass (74). Serum levels of sclerostin are positively associated with BMD and FM (75, 76). Sclerostin is regulated by estrogen and PTH (74). Postmenopausal women treated with PTH (1–34) have circulating levels of sclerostin decreased (77). Collectively, these findings suggest that the association between FM and BMD is likely mediated through sclerostin.

Distinguishing the role of LM vs FM as a determinant of BMD has clinical relevance. Physical activity is a strong determinant of bone mass acquisition during growth (78) and maintenance during postmenopause (79, 80) or late decades of life (81). A healthy dietary regime with rich vegetables is associated with healthy fat and bone accrual in children (82), but fast food and/or saturated fat intakes are adversely associated with bone health (83, 84). Thus, an association between BMD and LM suggests that an increase in physical activity may directly translate into protection against osteoporosis (85), whereas an association between BMD and FM implies that sex hormones and good nutrition may have a protective effect against bone loss (86). Results of this study seem to suggest that muscle strength and physical activity have a more prominent effect on bone health than hormones, particularly in men and premenopausal women.

As with any meta-analysis, exclusion of pertinent unpublished studies represents a threat to the validity of the result. However, in this analysis, we found no evidence of systematic publication bias by all methods (funnel, radial,

and standardized residual histogram). Nevertheless, there was a significant heterogeneity among the studies included in the analysis, and we dealt with this problem by a random-effects analysis and subgroup analysis (34). Possible heterogeneity expected in included studies could be due to the discrepancy of sample sizes and measurement of variables that were not the case in the random-effects analysis of correlation coefficients (87). Another threat of validity is that the association between LM, FM, and BMD might not be linear. However, in all original studies, the assumption of linearity appears to be tenable.

In conclusion, we have shown that whereas both LM and FM are associated with BMD, LM is more important than FM as a determinant of BMD in men and women of all ages and ethnicity. However, in postmenopausal women, the effects of LM and FM on BMD are equivalent. The importance of LM as a determinant of BMD underlines the concept that muscle mass or physical activity is an important component in the prevention of bone loss and osteoporosis in the general population. However, the significant association between FM and BMD suggests that sex hormones and nutrition also play an important role in the growth and maintenance of bone mass.

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